

# Geofeasibility scores

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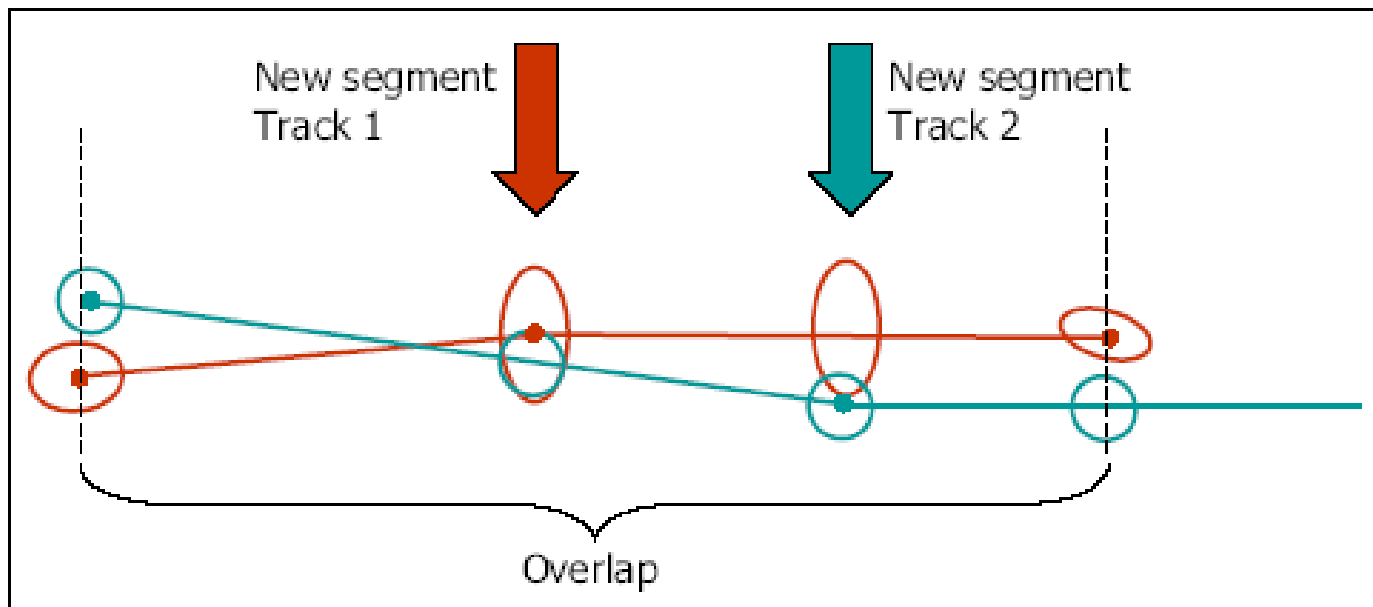
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# Problem Statement

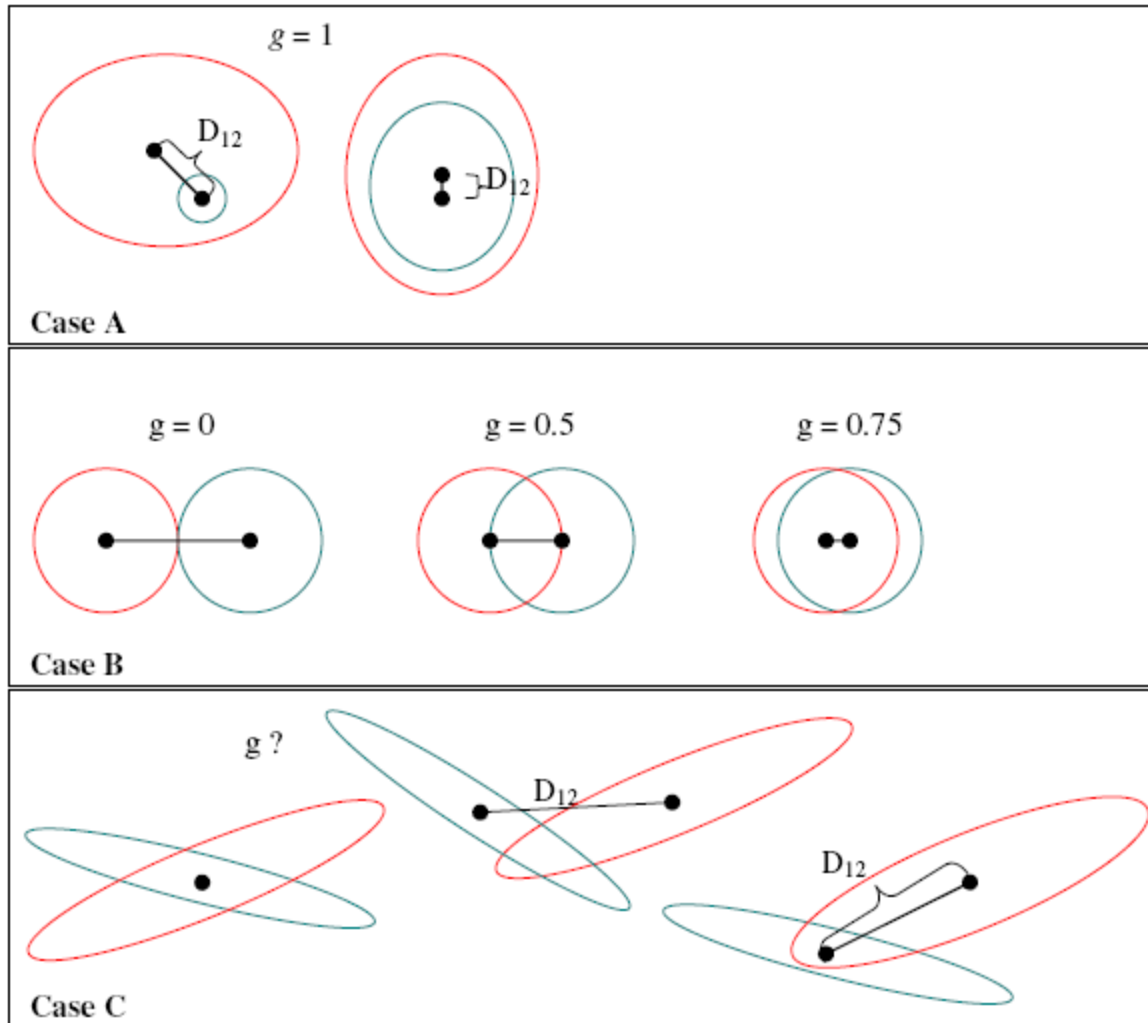
- For each report, the source associates an area of uncertainty (AOU) of elliptic shape delimiting a  $2\sigma$  probability area



# Problem Statement

- We need to define an optimal geo-feasibility score  $g$ , to quantify the overlap of two AOU.
- Here are three basic rules to define the function  $g$ :
- $g \in [0,1]$
- If ellipses do not touch,  $g = 0$
- If ellipses totally overlap,  $g = 1$

# Problem Statement



- The geo-feasibility score has to be suitable for very elongated ellipses as well as for circles
- Reasonable approximations are possible (e.g. approximating an ellipse to a circle is NOT a reasonable approximation)

# Candidate Metrics

- **The normalized area of overlap**
  - The area of overlap is normalized by the area of the smaller ellipse
  - Corresponds to human operator intuition.

# Candidate Metrics

- **Statistical Approaches**
- Integrated product of two Gaussian distributions
  - Corresponds to Bayes factor.
- Symmetric KL (Kullback-Leibler) divergence
  - Distance between two distributions.
- Generalized Likelihood Ratio (GLR)
  - Find most likely position of a single boat, evaluate the likelihood that it generated the reported distributions.

# Challenges

- Closed form analytical calculation of overlap area is not possible
- Numerical methods based on optimization are not fast enough

# Proposed solutions

- Newton's method to find intersection points and analytical approximation to the normalized area of the overlap
- Monte-Carlo integration to find the normalized area of overlap
- Generalized Likelihood Ratio gives a fast and meaningful approximation to the operator's intuition

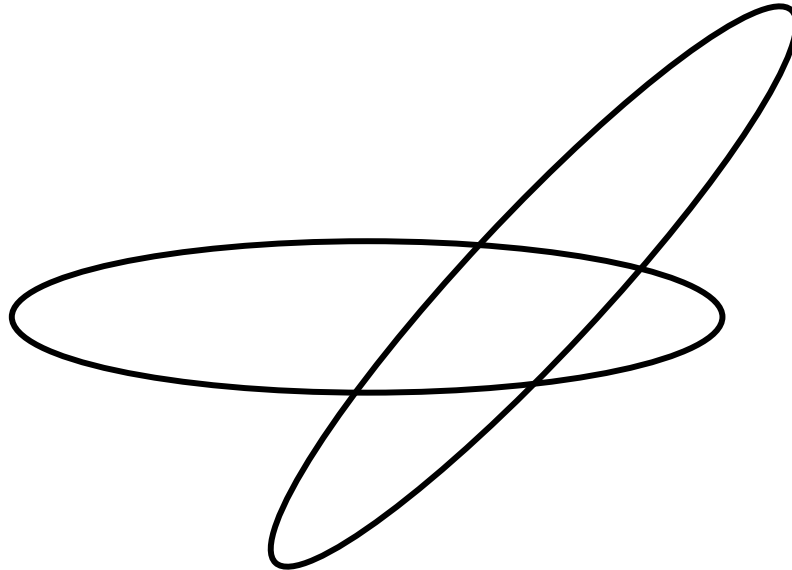
# Analytical Method

2 steps:

- Find the intersection points
  - Too hard to find analytically
- Calculate the area
  - Using integration in polar coordinates

# Finding points of Intersection

- Using Newton's Method
  - In-and-out method for starting points



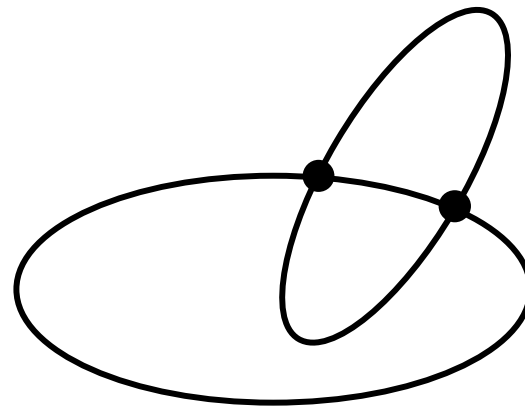
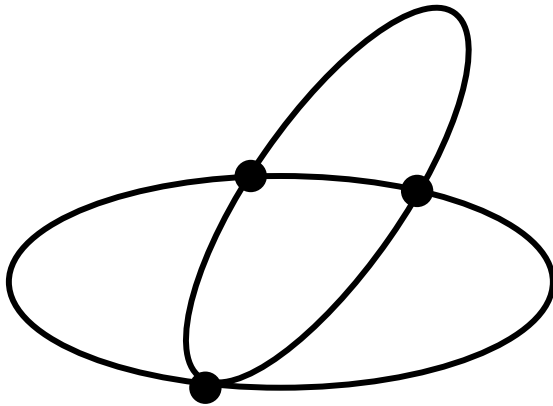
# Calculating area

3 cases:

- 0 or 1 points of intersection
  - Area is zero or is equal to the area of the smaller ellipse
- 2 or 3 points of intersection
- 4 points of intersection

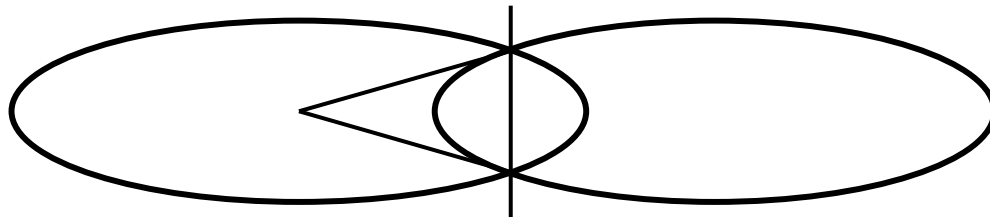
# 2 or 3 points of intersection

- Same case considering in-and-out technique

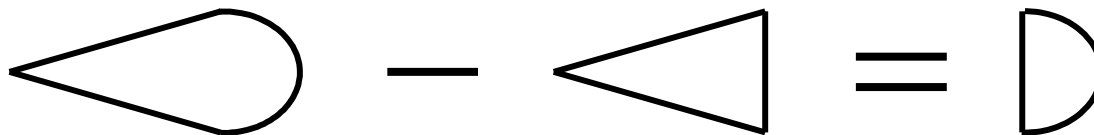


# For 2 points

- Using polar coordinates



Area of ellipse's portion – Area of the triangle



$$\text{D} + \text{C} = \text{Area of intersection}$$

# Calculation

- Ellipse in polar coordinates ( $R$  = radius)

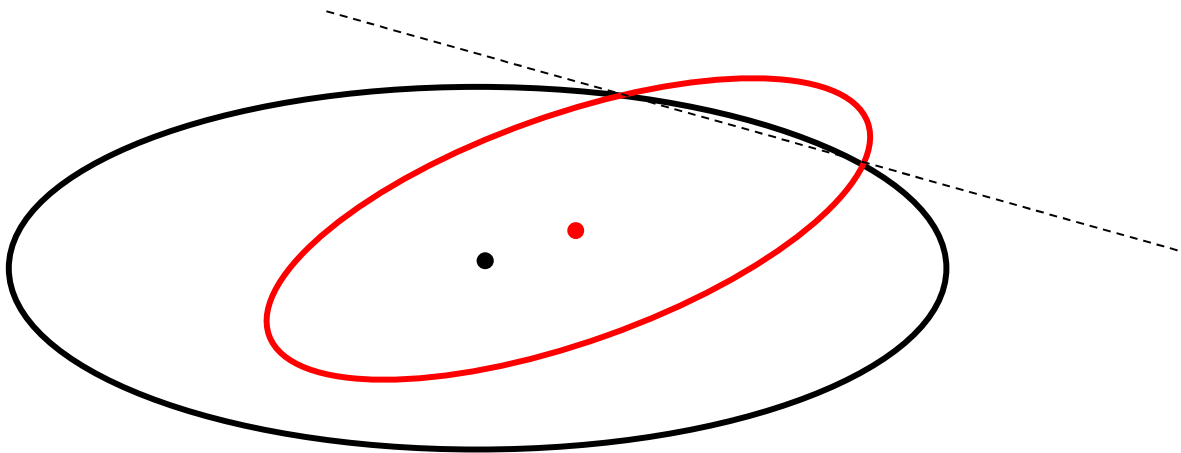
$$R(\theta) = \frac{ab}{\sqrt{a^2 + (b^2 - a^2)\cos^2(\theta)}}$$

- Integral

$$\int_{\theta_1}^{\theta_2} \frac{R^2(\theta)}{2} d\theta = \left[ \frac{ab}{2} \tan^{-1} \left( \frac{a \tan(\theta)}{b} \right) \right]_{\theta_1}^{\theta_2}$$

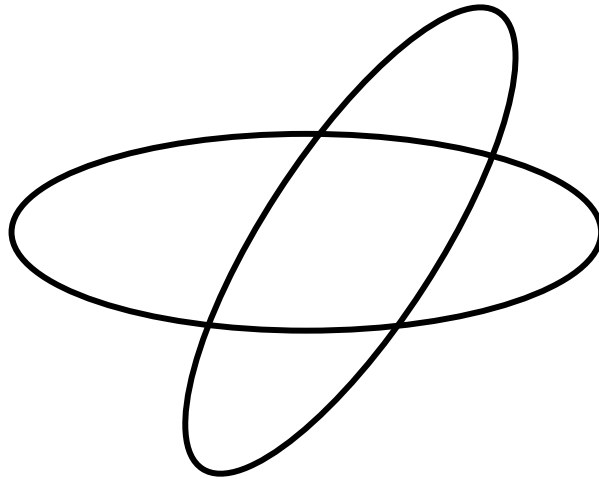
# Be careful...

- The integral uses inverse tangent function
  - $\theta \in \left(-\frac{\pi}{2}, \frac{\pi}{2}\right)$
- Too near ellipses center



# 4 points of intersection

- Extension of the 2 points case



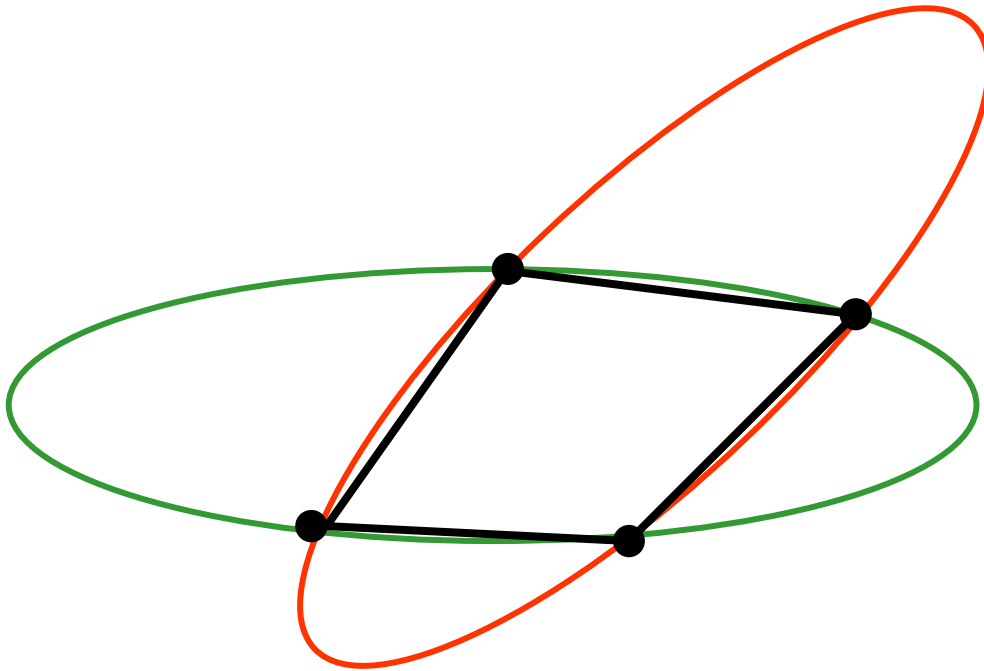
$E_i$  = Area of ellipse  $i$

$$a_i = \text{D}$$

$$\text{Area of intersection} = \frac{E_1 + E_2 - \sum_{i=1}^4 a_i}{2}$$

# Hard to code

- Approximation by a four sided object



# Performance

- The slow part of the program is if the ellipses actually intersect
- for 100,000 pairs of ellipses that intersect about 45% of the time, it takes between 640 and 710 seconds.

# Monte Carlo integration

- The algorithm computes an estimate of a multidimensional integral

$$I = \iiint_{x,y,z,\dots} f(x,y,z,\dots) dx dy dz \dots$$

- Naïve algorithm draws samples  $(x_i, y_i, z_i, \dots)$  uniformly from the integration area and estimates its value as follows

$$\hat{I} = \frac{1}{N} \sum_{i=1}^N f(x_i, y_i, z_i, \dots)$$

# Monte Carlo integration

- However, in many cases it is beneficial to draw samples from some pdf  $p(x,y,z,...)$

$$I = \iiint_{x,y,z,...} \frac{f(x,y,z,...)}{p(x,y,z,...)} p(x,y,z,...) dx dy dz ...$$

- And use the so called **importance sampling**

$$\hat{I} = \frac{1}{N} \sum_{i=1}^N \frac{f(x_i, y_i, z_i, ...)}{p(x_i, y_i, z_i, ...)}$$

# Monte Carlo integration

- **Normalized intersection area** of two ellipses defined by the regions  $S_1$  and  $S_2$

$$g_a = \frac{\int 1_{\mathbf{x} \in S_1 \cap S_2} d\mathbf{x}}{\int 1_{\mathbf{x} \in \min(S_1, S_2)} d\mathbf{x}}$$

- Can be estimated using Monte Carlo with importance sampling

$$g_a = \frac{\sum_{i=1}^N \frac{1}{p(\mathbf{x}_i)} 1_{\mathbf{x}_i \in S_1 \cap S_2}}{\sum_{i=1}^N \frac{1}{p(\mathbf{x}_i)} 1_{\mathbf{x}_i \in \min(S_1, S_2)}}$$

# Monte Carlo integration

- **Integrated product of two Gaussian distributions**

$$g_p = \int_{\mathbf{x}} p_1(\mathbf{x} | \Theta_1) p_2(\mathbf{x} | \Theta_2) d\mathbf{x}$$

- Can be estimated using Monte Carlo even without importance sampling

$$\hat{g}_p = \frac{1}{N} \sum_{i=1}^{N/2} p_1(\mathbf{x}_i^2 | \Theta_1) + \frac{1}{N} \sum_{i=1}^{N/2} p_2(\mathbf{x}_i^1 | \Theta_2)$$

# Generalized Likelihood Ratio

- We assume that a source provides us with measurements of target positions and Maximum Likelihood (ML) estimators of covariance matrices

$$\mathbf{y}_1, \quad \hat{\mathbf{R}}_1, \quad \mathbf{y}_2, \quad \hat{\mathbf{R}}_2$$

- And we choose to construct a test statistic to discriminate between the two hypotheses

$$H_1 : \quad \boldsymbol{\mu}_1 = \boldsymbol{\mu}_2$$

$$H_0 : \quad \boldsymbol{\mu}_1 \neq \boldsymbol{\mu}_2$$

# Generalized Likelihood Ratio

- In this case Uniformly Most Powerful test does not exist.
- However, we can resort to a suboptimum statistic that is called GLR

$$\Lambda(\mathbf{y}_1, \mathbf{y}_2) = \frac{\max_{\Theta} p(\mathbf{y}_1, \mathbf{y}_2 \mid \Theta, H_1)}{\max_{\Theta} p(\mathbf{y}_1, \mathbf{y}_2 \mid \Theta, H_0)}$$

- Here  $\Theta$  stands for all the unknown parameters

# Generalized Likelihood Ratio

- Given the Gaussian and independence assumptions we have

$$p(\mathbf{y}_1, \mathbf{y}_2 \mid \Theta, H) = \frac{1}{2\pi \sqrt{\det(\mathbf{R}_1) \det(\mathbf{R}_2)}} \exp \left[ -\frac{1}{2} (\mathbf{y}_1 - \boldsymbol{\mu}_1)^T \mathbf{R}_1^{-1} (\mathbf{y}_1 - \boldsymbol{\mu}_1) - \frac{1}{2} (\mathbf{y}_2 - \boldsymbol{\mu}_2)^T \mathbf{R}_2^{-1} (\mathbf{y}_2 - \boldsymbol{\mu}_2) \right]$$

- We deduce immediately that

$$\max_{\Theta} p(\mathbf{y}_1, \mathbf{y}_2 \mid \Theta, H_0) = \frac{1}{2\pi \sqrt{\det(\hat{\mathbf{R}}_1) \det(\hat{\mathbf{R}}_2)}}$$

# Generalized Likelihood Ratio

- To find the ML estimator of the mean under  $H_1$  we use  $\mu_1 = \mu_2 = \mu$  and equate partial derivative to 0:

$$\frac{\partial}{\partial \boldsymbol{\mu}} \ln p(\mathbf{y}_1, \mathbf{y}_2 \mid \Theta, H_1) = \mathbf{R}_1^{-1}(\mathbf{y}_1 - \boldsymbol{\mu}) + \mathbf{R}_2^{-1}(\mathbf{y}_2 - \boldsymbol{\mu}) = \bar{\mathbf{0}}$$

- Which results in the following expression for the ML estimator of the mean

$$\hat{\boldsymbol{\mu}} = \left[ \mathbf{R}_1^{-1} + \mathbf{R}_2^{-1} \right]^{-1} \left[ \mathbf{R}_1^{-1} \mathbf{y}_1 + \mathbf{R}_2^{-1} \mathbf{y}_2 \right]$$

# Generalized Likelihood Ratio

- Substituting this estimator into likelihood ratio we get statistic of the form

$$\Lambda(\mathbf{y}_1, \mathbf{y}_2) = \exp\left[-\frac{1}{2}\Delta^T \hat{\mathbf{R}}_{\Delta}^{-1}\Delta\right]$$

- Where  $\Delta$  is the difference between measurements:

$$\Delta = \mathbf{y}_1 - \mathbf{y}_2$$

# Generalized Likelihood Ratio

- And the estimate of the covariance inverse has the following “nice” expression

$$\begin{aligned}\hat{\mathbf{R}}_{\Delta}^{-1} = & \hat{\mathbf{R}}_2^{-1} \left[ \hat{\mathbf{R}}_1^{-1} + \hat{\mathbf{R}}_2^{-1} \right]^{-1} \hat{\mathbf{R}}_1^{-1} \left[ \hat{\mathbf{R}}_1^{-1} + \hat{\mathbf{R}}_2^{-1} \right]^{-1} \hat{\mathbf{R}}_2^{-1} + \\ & + \hat{\mathbf{R}}_1^{-1} \left[ \hat{\mathbf{R}}_1^{-1} + \hat{\mathbf{R}}_2^{-1} \right]^{-1} \hat{\mathbf{R}}_2^{-1} \left[ \hat{\mathbf{R}}_1^{-1} + \hat{\mathbf{R}}_2^{-1} \right]^{-1} \hat{\mathbf{R}}_1^{-1}\end{aligned}$$

- However, using the rule “the inverse of the product is equal to the product of inverses in reversed order” we can show that

$$\hat{\mathbf{R}}_{\Delta}^{-1} = \left[ \hat{\mathbf{R}}_1 + \hat{\mathbf{R}}_2 \right]^{-1}$$

# Generalized Likelihood Ratio

- Thus the geofeasibility score based on GLR admits the following simple and intuitive form

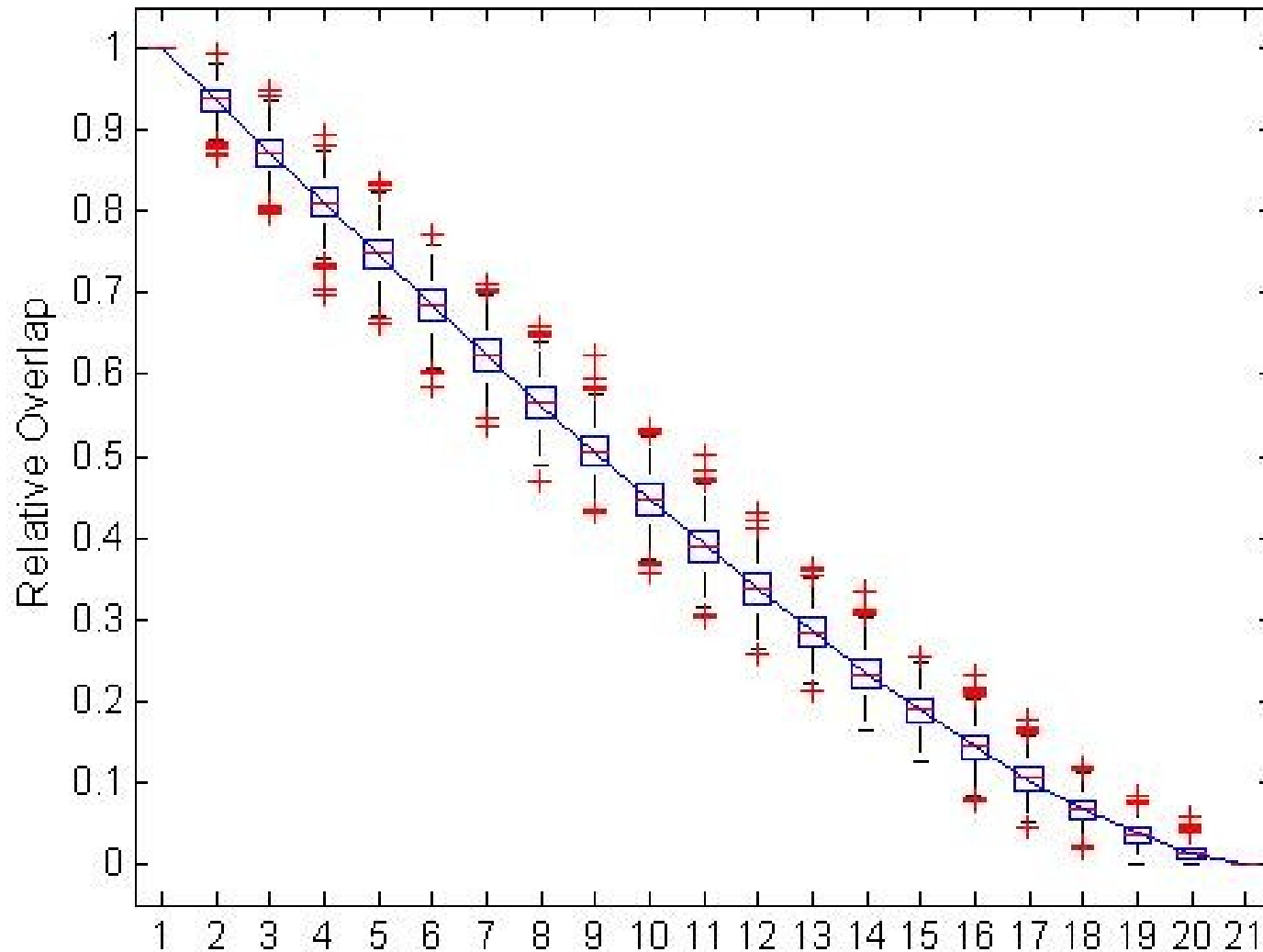
$$g_G \equiv \Lambda(\mathbf{y}_1, \mathbf{y}_2)$$

$$g_G = \exp \left[ -\frac{1}{2} \Delta^T \left[ \hat{\mathbf{R}}_1 + \hat{\mathbf{R}}_2 \right]^{-1} \Delta \right]$$

# Simulation results

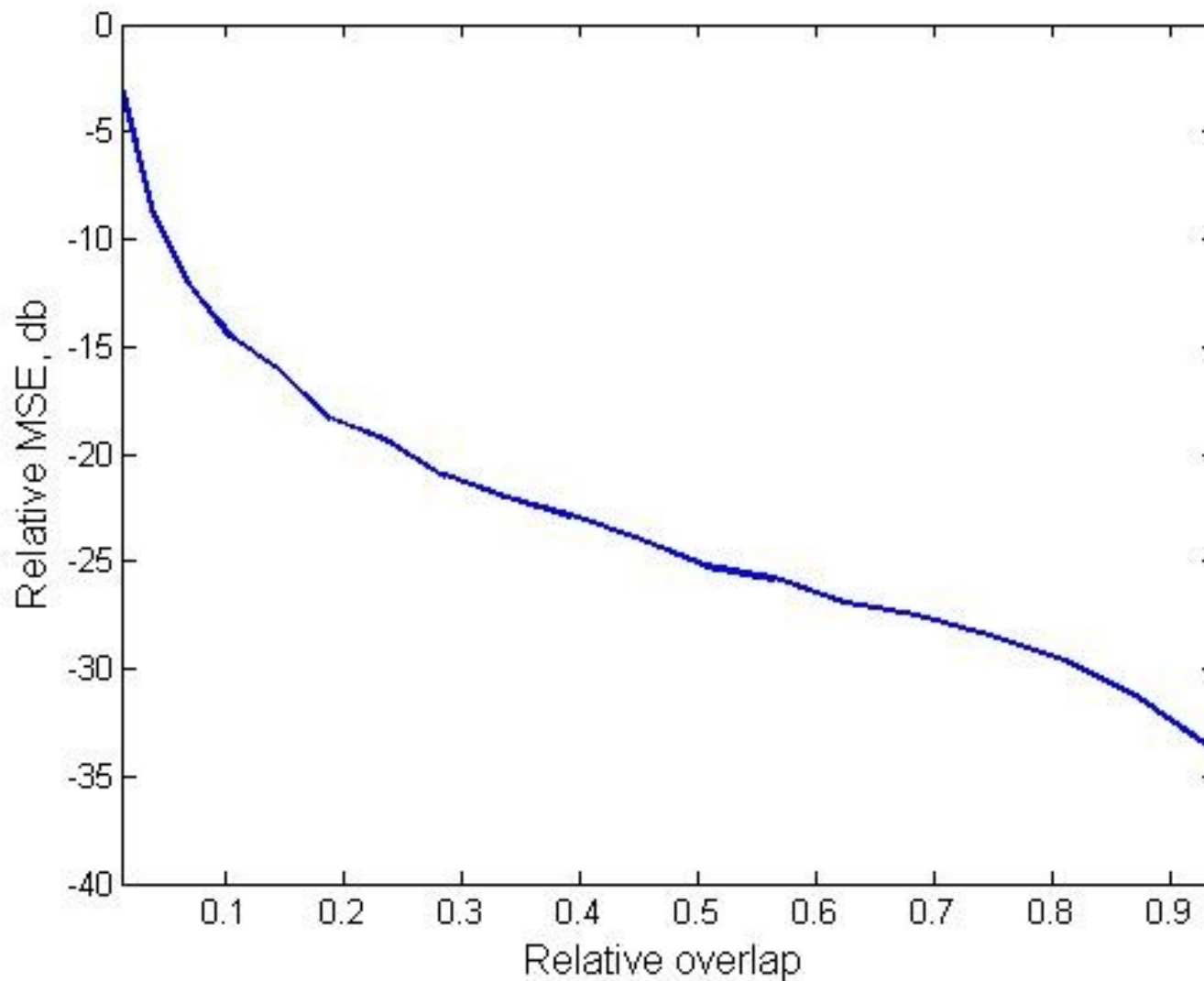
Error bars, Monte-Carlo integration of the overlap area

$N=500$



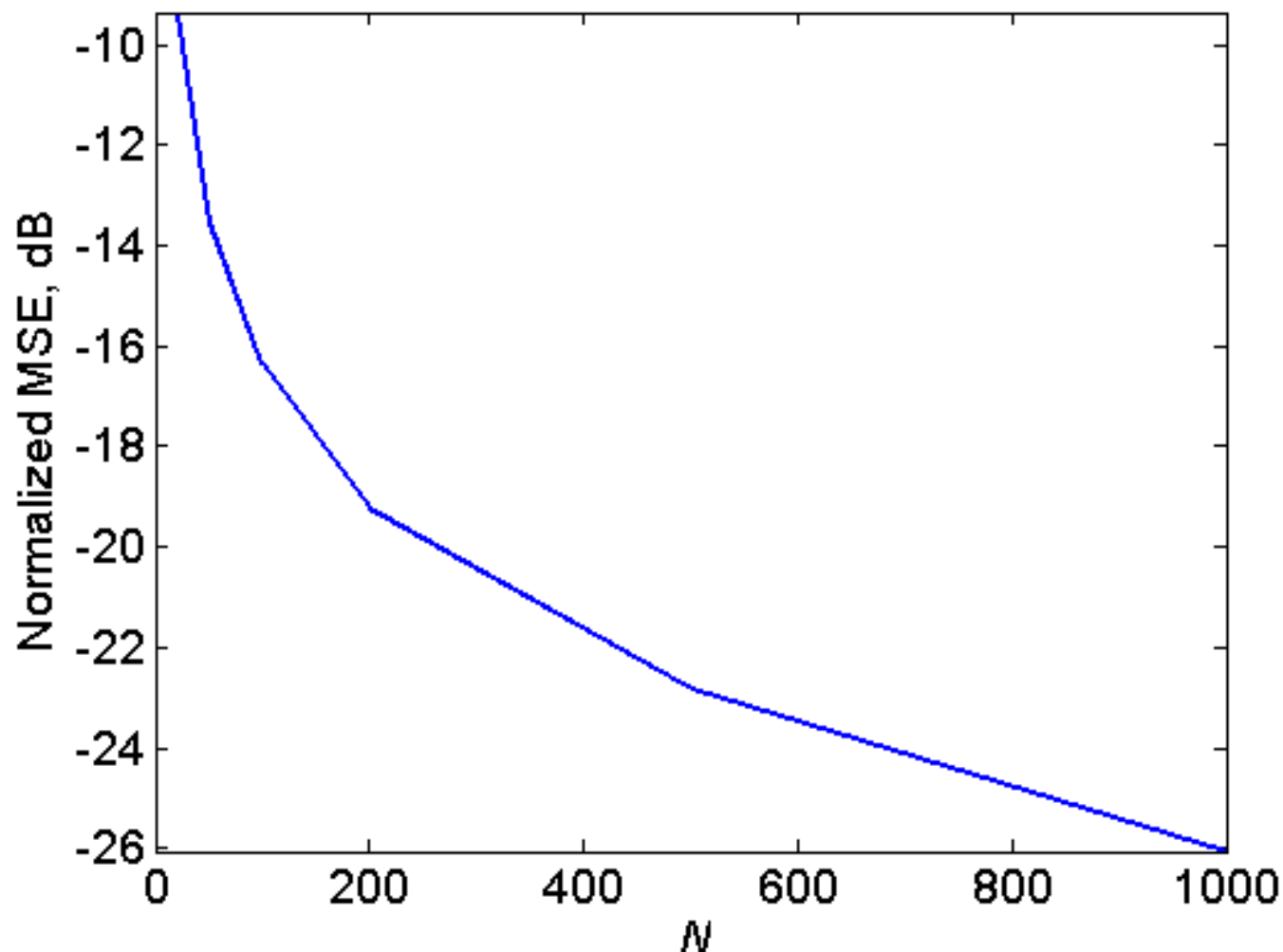
# Simulation results

Monte Carlo Mean Squared Error,  $N=500$



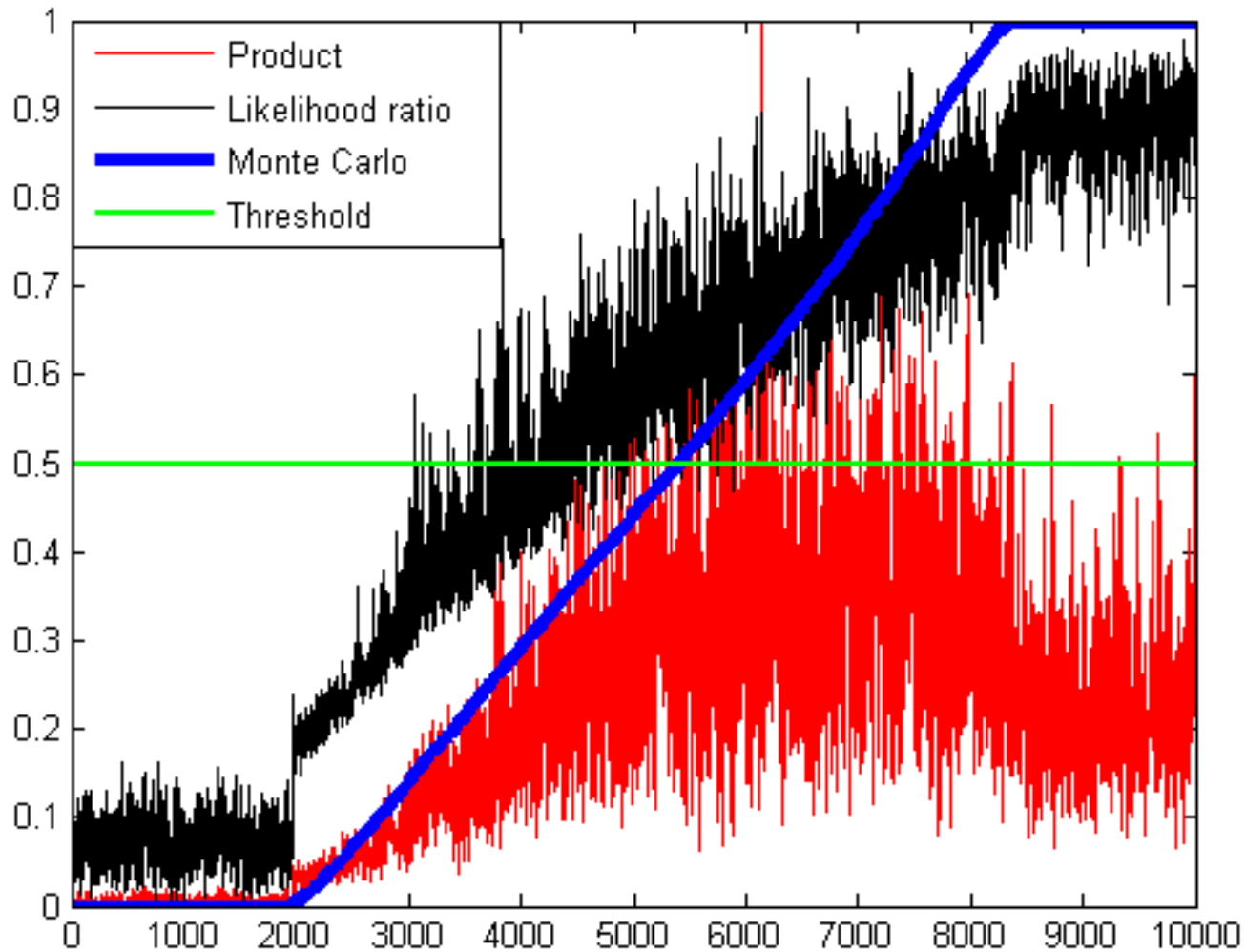
# Simulation results

Monte-Carlo error for a fixed value of relative overlap area equal to 0.391



# Simulation results

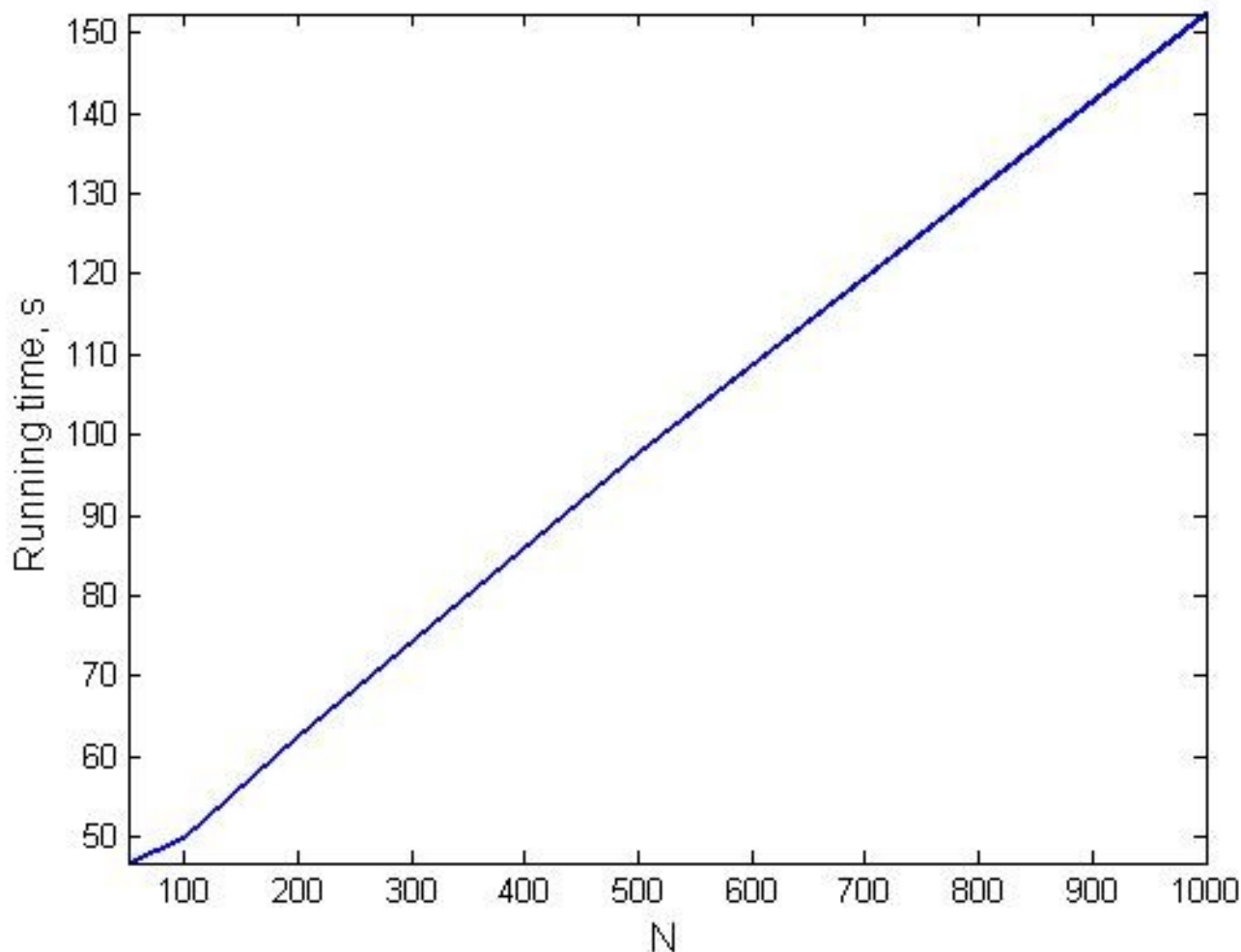
## Comparison of statistics



# Simulation results

Calculation time for 100,000 evaluations, Monte-Carlo.

**Calculation time for GLR is 4 seconds  
(0.04 ms/evaluation)**



# Questions